

# Approximating EPT Densities via RARL2 with Financial Modelling Applications of two sided EPT Densities

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# INTRODUCTION

# EPT Functions on Half Line

$$f(x) = \operatorname{Re} \left( \sum_{k=1}^K p_k(x) e^{\mu_k x} \right) \quad x \geq 0$$

$$f(x) = \mathbf{c} e^{\mathbf{A}x} \mathbf{b} \quad \text{if } x \geq 0$$

$\mathbf{A}$  is an  $(n \times n)$  matrix,  $\mathbf{c}$  a  $(1 \times n)$  row vector and  $\mathbf{b}$  a  $(n \times 1)$  column vector

- Class contains the real polynomials, real exponential functions and scaled real trigonometric polynomials
- Closed under addition and multiplication
- Strictly proper rational characteristic function

$$\phi(iu) = \int_0^{\infty} e^{iux} \mathbf{c} e^{\mathbf{A}x} \mathbf{b} dx = -\mathbf{c}(\mathbf{I}iu + \mathbf{A})^{-1} \mathbf{b} = \frac{p(iu)}{q(iu)}$$

$q$  is a polynomial of degree  $n$  while  $p$  is a polynomial of degree  $m < n$

$\phi$  a rational function of McMillan degree  $n$

# EPT Density

Triple  $(\mathbf{A}, \mathbf{b}, \mathbf{c})$  such that

$$f(x) = \mathbf{c}e^{\mathbf{A}x}\mathbf{b} \geq 0 \quad \forall x \geq 0$$

- Test for non-negativity, Hanzon and Holland (2010)
- We use Minimal Realizations (use a minimal realization algorithm if necessary)
- Stability of  $\mathbf{A}$  as eigenvalues located in open left plane

$$\lim_{x \rightarrow \infty} f(x) = \lim_{x \rightarrow \infty} \mathbf{c}e^{\mathbf{A}x}\mathbf{b} = 0$$

- Normalization due to finite integrability

$$\int_0^{\infty} \mathbf{c}e^{\mathbf{A}x}\mathbf{b}dx = -\mathbf{c}\mathbf{A}^{-1}\mathbf{b} < \infty$$

Triple  $(\mathbf{A}, \mathbf{b}, \tilde{\mathbf{c}})$  now represents a probability density function

$$\tilde{\mathbf{c}} = \frac{\mathbf{c}}{-\mathbf{c}\mathbf{A}^{-1}\mathbf{b}}$$

## 2-EPT Density

A 2-EPT density requires two EPT densities defined on either half real line

$$f(x) = \begin{cases} \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N & \text{if } x \leq 0 \\ \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P & \text{if } x > 0 \end{cases} \quad (1)$$

$\mathbf{A}_N$  is “anti-stable” with all eigenvalues located in open right half plane

$$\lim_{x \rightarrow \infty} \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P = \lim_{x \rightarrow -\infty} \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N = 0$$

Finite integrability allows for normalization

$$\int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^0 \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N dx + \int_0^{\infty} \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P dx = \mathbf{c}_N \mathbf{A}_N^{-1} \mathbf{b}_N + \mathbf{c}_P \mathbf{A}_P^{-1} \mathbf{b}_P$$

Triples  $(\mathbf{A}_P, \mathbf{b}_P, \tilde{\mathbf{c}}_P)$  and  $(\mathbf{A}_N, \mathbf{b}_N, \tilde{\mathbf{c}}_N)$  define probability density function where  $\tilde{\mathbf{c}}_P$  and  $\tilde{\mathbf{c}}_N$  are scaled ensure density integrates to unity.

# State Space Representation

Rational Characteristic function decomposed

$$\begin{aligned}
 \phi(iu) &= \int_{-\infty}^0 \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N e^{iu x} dx + \int_0^{\infty} \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P e^{iu x} dx \\
 &= \mathbf{c}_N (\mathbf{I}iu + \mathbf{A}_N)^{-1} \mathbf{b}_N - \mathbf{c}_P (\mathbf{I}iu + \mathbf{A}_P)^{-1} \mathbf{b}_P \\
 &= \phi_N(iu) + \phi_P(iu) = \frac{p(iu)}{q(iu)}
 \end{aligned} \tag{2}$$

Represented in state space form

$$\phi(u) = \pi \left( \begin{array}{cc|c} \mathbf{A}_P & 0 & \mathbf{b}_P \\ 0 & \mathbf{A}_N & \mathbf{b}_N \\ \hline -\mathbf{c}_P & \mathbf{c}_N & 0 \end{array} \right) \tag{3}$$

Numerous operations can be carried out when in this form using state space techniques of Hanzon and Ober (2001), Hanzon and Scherrer (2008)

# Point Mass at Zero

A natural generalisation of the EPT class also allows for a pointmass at zero

$$f(x) + d\delta \quad 0 \leq d \leq 1$$

where  $\delta$  is the delta distribution and  $f(x)$  a 2-EPT density function

Hence the state space representation is given by

$$\phi(u) = \pi \left( \begin{array}{cc|c} \mathbf{A}_P & 0 & \mathbf{b}_P \\ 0 & \mathbf{A}_N & \mathbf{b}_N \\ \hline -\mathbf{c}_P & \mathbf{c}_N & d \end{array} \right)$$

which is the rational characteristic function

$$\phi(u) = \mathbf{c}_N (\mathbf{l}iu + \mathbf{A}_N)^{-1} \mathbf{b}_N - \mathbf{c}_P (\mathbf{l}iu + \mathbf{A}_P)^{-1} \mathbf{b}_P + d$$

# Phase-Type, Matrix Exponential and Positive Realizations

Positive Realizations are strict subset of EPT class

- Non-negativity guaranteed
- Positive realization may not exist
- If positive realization does exist, dimensions may be greater than McMillan degree

Similar arguments apply to phase type distributions where a Phase-Type distribution may not exist for a given characteristic function.

Matrix Exponential class is equivalent to EPT class

- Realization could be found from rational characteristic function using companion matrix
- No reference to minimal realizations in literature
- Matrix Exponential only defined on non-negative half line



## OPERATIONS

# Scaling

Let  $X$  be a 2-EPT random variable with density function given by Eq. (1). Let  $\alpha \in \mathbb{R}^1$  such that  $\alpha > 0$  and  $Y = \alpha X$ .

$$\begin{aligned}
 G(y_0) &= \mathbb{P}(Y \leq y_0) \\
 &= \mathbb{P}(\alpha X \leq y_0) \\
 &= \mathbb{P}(X \leq \frac{y_0}{\alpha}) \\
 &= F(\frac{y_0}{\alpha})
 \end{aligned}$$

Differentiating we get

$$g(y_0) = f(\frac{y_0}{\alpha}) \frac{1}{\alpha}$$

and leaving  $x_0 = \frac{y_0}{\alpha}$  we obtain that

$$\alpha g(\alpha x_0) = f(x_0)$$

Hence if our original triples for  $X$  were  $(\mathbf{A}_P, \mathbf{b}_P, \mathbf{c}_P)$  and  $(\mathbf{A}_N, \mathbf{b}_N, \mathbf{c}_N)$  the associated triples for  $Y$  are given by  $(\alpha \mathbf{A}_P, \mathbf{b}_P, \alpha \mathbf{c}_P)$  and  $(\alpha \mathbf{A}_N, \mathbf{b}_N, \alpha \mathbf{c}_N)$

# Translation

Let  $X$  be a random variable with a 2-EPT density function,  $f$ , as in Eq. (1).  
 Let  $x_0 \in \mathbb{R}^1$ . The random variable  $X + x_0$  has a density function  $g(x)$  such that  
 $g(x) = f(x - x_0)$ .

$$g(x) = f(x - x_0) = \begin{cases} \mathbf{c}_N e^{\mathbf{A}_N x} e^{-\mathbf{A}_N x_0} \mathbf{b}_N & \text{if } x \leq x_0 \\ \mathbf{c}_P e^{\mathbf{A}_P x} e^{-\mathbf{A}_P x_0} \mathbf{b}_P & \text{if } x > x_0 \end{cases}$$

$$\begin{aligned} \phi(iu) &= \int_{-\infty}^{x_0} \mathbf{c}_N e^{\mathbf{A}_N x} e^{-\mathbf{A}_N x_0} e^{iu x} \mathbf{b}_N dx + \int_{x_0}^{\infty} \mathbf{c}_P e^{\mathbf{A}_P x} e^{-\mathbf{A}_P x_0} e^{iu x} \mathbf{b}_P dx \\ &= \mathbf{c}_N (\mathbf{A}_N + u\mathbf{i})^{-1} e^{(\mathbf{A}_N + u\mathbf{i})x_0} e^{-\mathbf{A}_N x_0} \mathbf{b}_N - \mathbf{c}_P (\mathbf{A}_P + u\mathbf{i})^{-1} e^{(\mathbf{A}_P + u\mathbf{i})x_0} e^{-\mathbf{A}_P x_0} \mathbf{b}_P \end{aligned}$$

which is not rational.

We do not consider this operation further as it does not fit into the EPT class

## Sums of 2-EPT Functions

Consider the mixture density  $g(x)$  of two 2-EPT density functions  $f_1(x)$  and  $f_2(x)$  as defined in Eq. (1) with weights  $w$  and  $(1 - w)$  resp. where  $0 < w < 1$ .

$$g(x) = \begin{cases} w \mathbf{c}_{N,1} e^{\mathbf{A}_{N,1} x} \mathbf{b}_{N,1} + (1 - w) \mathbf{c}_{N,2} e^{\mathbf{A}_{N,2} x} \mathbf{b}_{N,2} = \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N & \text{if } x \leq 0 \\ w \mathbf{c}_{P,1} e^{\mathbf{A}_{P,1} x} \mathbf{b}_{P,1} + (1 - w) \mathbf{c}_{P,2} e^{\mathbf{A}_{P,2} x} \mathbf{b}_{P,2} = \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P & \text{if } x > 0 \end{cases}$$

For  $x > 0$  we have  $\mathbf{c}_P = (w \mathbf{c}_{P,1}, (1 - w) \mathbf{c}_{P,2})$ ,  $\mathbf{b}_P = (\mathbf{b}_{P,1}; \mathbf{b}_{P,2})$

We split the  $\mathbf{A}$  matrices into a block diagonal structure as follows

$$\mathbf{A}_N = \begin{pmatrix} \mathbf{A}_{N,1} & 0 \\ 0 & \mathbf{A}_{N,2} \end{pmatrix}, \quad \mathbf{A}_P = \begin{pmatrix} \mathbf{A}_{P,1} & 0 \\ 0 & \mathbf{A}_{P,2} \end{pmatrix}$$

If we allow  $w < 0$  then we must test for non-negativity

## Product of two 2-EPT Densities

To calculate the product of two 2-EPT probability densities  $f_1(x)$  and  $f_2(x)$

$$f_i(x) = \begin{cases} \mathbf{c}_{N,i} e^{\mathbf{A}_{N,i} x} \mathbf{b}_{N,i} & \text{if } x \leq 0 \\ \mathbf{c}_{P,i} e^{\mathbf{A}_{P,i} x} \mathbf{b}_{P,i} & \text{if } x > 0 \end{cases}$$

where  $\mathbf{A}_{N,i}$  and  $\mathbf{A}_{P,i}$  are square matrices of size  $n_{N,i}$  and  $n_{P,i}$  respectively.

$$g(x) = f_1(x) f_2(x) = \begin{cases} \mathbf{c}_{N,1} e^{\mathbf{A}_{N,1} x} \mathbf{b}_{N,1} \mathbf{c}_{N,2} e^{\mathbf{A}_{N,2} x} \mathbf{b}_{N,2} = \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N & \text{if } x \leq 0 \\ \mathbf{c}_{P,1} e^{\mathbf{A}_{P,1} x} \mathbf{b}_{P,1} \mathbf{c}_{P,2} e^{\mathbf{A}_{P,2} x} \mathbf{b}_{P,2} = \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P & \text{if } x > 0 \end{cases}$$

For  $x > 0$  we obtain the triple  $(\mathbf{A}_P, \mathbf{b}_P, \mathbf{c}_P)$

$$\mathbf{A}_P = \mathbf{A}_{P,1} \otimes \mathbf{I}_{n_{P,2}} + \mathbf{I}_{n_{P,1}} \otimes \mathbf{A}_{P,2}$$

$$\mathbf{b}_P = \mathbf{b}_{P,1} \otimes \mathbf{b}_{P,2}$$

$$\mathbf{c}_P = \mathbf{c}_{P,1} \otimes \mathbf{c}_{P,2}$$

where  $\mathbf{I}_n$  is the  $n \times n$  identity matrix.

# Moment Calculations via Characteristic Function

The characteristic function of a 2-EPT density is

$$\phi(u) = \mathbf{c}_N (\mathbf{I}u + \mathbf{A}_N)^{-1} \mathbf{b}_N - \mathbf{c}_P (\mathbf{I}u + \mathbf{A}_P)^{-1} \mathbf{b}_P$$

The  $k^{\text{th}}$  derivative of  $\phi$  is again a rational functions with the same pole locations as  $\phi$

$$\frac{d^k \phi(u)}{du^k} = (-1)^k k! \mathbf{c}_N (\mathbf{I}u + \mathbf{A}_N)^{-(k+1)} \mathbf{b}_N - (-1)^k k! \mathbf{c}_P (\mathbf{I}u + \mathbf{A}_P)^{-(k+1)} \mathbf{b}_P$$

Evaluated at  $u = 0$  we get

$$\left. \frac{d^k \phi(u)}{du^k} \right|_{u=0} = (-1)^k k! \mathbf{c}_N (\mathbf{A}_N)^{-(k+1)} \mathbf{b}_N - (-1)^k k! \mathbf{c}_P (\mathbf{A}_P)^{-(k+1)} \mathbf{b}_P$$

All moments must exist

## Calculating Moments using Direct Integration

Calculating the  $k^{\text{th}}$  moment by integrating  $x^k f(x)$  over the relevant domain. It must be noted that  $p(x) = x^k = \mathbf{c}_k e^{\mathbf{A}_k x} \mathbf{b}_k$  where  $\mathbf{A}_k$  is a  $(k+1)$  square nilpotent matrix.

$$\begin{aligned}
 \int_{-d}^d x^k f(x) dx &= \int_{-d}^0 \mathbf{c}_k e^{\mathbf{A}_k x} \mathbf{b}_k \mathbf{c}_N e^{\mathbf{A}_N x} \mathbf{b}_N dx + \int_0^d \mathbf{c}_k e^{\mathbf{A}_k x} \mathbf{b}_k \mathbf{c}_P e^{\mathbf{A}_P x} \mathbf{b}_P dx \\
 &= \int_{-d}^0 \tilde{\mathbf{c}}_N e^{\tilde{\mathbf{A}}_N x} \tilde{\mathbf{b}}_N dx + \int_0^d \tilde{\mathbf{c}}_P e^{\tilde{\mathbf{A}}_P x} \tilde{\mathbf{b}}_P dx \\
 &= \tilde{\mathbf{c}}_N \tilde{\mathbf{A}}_N^{-1} \tilde{\mathbf{b}}_N - \tilde{\mathbf{c}}_N \tilde{\mathbf{A}}_N^{-1} e^{-\tilde{\mathbf{A}}_N d} \tilde{\mathbf{b}}_N + \tilde{\mathbf{c}}_P \tilde{\mathbf{A}}_P^{-1} e^{\tilde{\mathbf{A}}_P d} \tilde{\mathbf{b}}_P - \tilde{\mathbf{c}}_P \tilde{\mathbf{A}}_P^{-1} \tilde{\mathbf{b}}_P
 \end{aligned}$$

where the triples  $(\tilde{\mathbf{A}}_P, \tilde{\mathbf{b}}_P, \tilde{\mathbf{c}}_P)$  and  $(\tilde{\mathbf{A}}_N, \tilde{\mathbf{b}}_N, \tilde{\mathbf{c}}_N)$  can be determined using the product formulae

In the limit  $d \rightarrow \infty$  we obtain

$$\int_{-\infty}^{\infty} x^k \mathbf{c} e^{\mathbf{A} x} \mathbf{b} dx = \tilde{\mathbf{c}}_N \tilde{\mathbf{A}}_N^{-1} \tilde{\mathbf{b}}_N - \tilde{\mathbf{c}}_P \tilde{\mathbf{A}}_P^{-1} \tilde{\mathbf{b}}_P$$

## Minimum of 2-EPT Random Variables

Consider two independent 2-EPT random variables  $X, Y$  with densities  $f_i$  for  $i = 1, 2$ . The following Cumulative Distribution Functions (CDF) are derived

$$F_i(z) = \begin{cases} \mathbf{c}_{N,i} \mathbf{A}_{N,i}^{-1} e^{\mathbf{A}_{N,i} z} \mathbf{b}_{N,i} & \text{if } z \leq 0 \\ \mathbf{c}_{P,i} \mathbf{A}_{P,i}^{-1} e^{\mathbf{A}_{P,i} z} \mathbf{b}_{P,i} + 1 & \text{if } z > 0 \end{cases}$$

If we let  $W = \text{Min}(X, Y)$ , using the CDFs and the property

$$\mathbb{P}[W \leq w] = 1 - \mathbb{P}[X \geq w] \mathbb{P}[Y \geq w] \quad \dots \quad \text{by independence}$$

we obtain the density function  $h(w)$  for the minimum

$$h(w) = \begin{cases} \mathbf{c}_{N,1} e^{\mathbf{A}_{N,1} w} \mathbf{b}_{N,1} + \mathbf{c}_{N,2} e^{\mathbf{A}_{N,2} w} \mathbf{b}_{N,2} - \mathbf{c}_N \tilde{\mathbf{A}}_{N,1,N,2}^{-1} \mathbf{A}_{N,1,N,2} e^{\tilde{\mathbf{A}}_{N,1,N,2} w} \mathbf{b}_N & w \leq 0 \\ -\mathbf{c}_P \tilde{\mathbf{A}}_{P,1,P,2}^{-1} \mathbf{A}_{P,1,P,2} e^{\tilde{\mathbf{A}}_{P,1,P,2} w} \mathbf{b}_P & w > 0 \end{cases}$$

A similar result can be found for  $Z = \text{Max}(X, Y)$



## Convolution of EPT Densities

Given two independent 2-EPT random variables  $X, Y$  with triples  $(\mathbf{A}_1, \mathbf{b}_1, \mathbf{c}_1)$  and  $(\mathbf{A}_2, \mathbf{b}_2, \mathbf{c}_2)$  respectively. The state space representation of  $Z = X + Y$  is

$$\phi_Z(u) = \phi_X(u)\phi_Y(u) = \pi \left( \begin{array}{cc|c} \mathbf{A}_1 & \mathbf{b}_1 \mathbf{c}_2 & \mathbf{b}_2 \\ 0 & \mathbf{A}_2 & 0 \\ \hline \mathbf{c}_1 & 0 & 0 \end{array} \right) = \pi \left( \begin{array}{c|c} \mathbf{A} & \mathbf{b} \\ \hline \mathbf{c} & 0 \end{array} \right) \quad (4)$$

where

$$\mathbf{A}_1 = \begin{pmatrix} \mathbf{A}_{P,1} & 0 \\ 0 & \mathbf{A}_{N,1} \end{pmatrix}, \quad \mathbf{A}_2 = \begin{pmatrix} \mathbf{A}_{P,2} & 0 \\ 0 & \mathbf{A}_{N,2} \end{pmatrix}$$

To ensure the rational characteristic function of  $Z$  in Eq. (4) can be decomposed appropriately we must block diagonalise  $\mathbf{A}$  and separate the eigenvalues according to the sign of their real part

## Block Diagonalise $\mathbf{A}$ while separating the Eigenvalues

A basis transformation is used to separate the eigenvalues of  $\mathbf{A}$ . Assuming the sequence of transformations can be expressed by  $\mathbf{Q}$  we obtain

$$\phi_Z(u) = \pi \left( \frac{\mathbf{QAQ}^{-1} \mid \mathbf{Qb}}{\mathbf{cQ}^{-1} \mid 0} \right), \quad \text{where } \mathbf{QAQ}^{-1} = \begin{pmatrix} \mathcal{A}_P & \mathcal{V} \\ 0 & \mathcal{A}_N \end{pmatrix}$$

A transformation with  $T$  is used to block diagonalise  $\mathbf{QAQ}^{-1}$

$$\mathbf{TQAQ}^{-1}\mathbf{T}^{-1} = \begin{pmatrix} \mathcal{A}_P & \mathcal{T}\mathcal{A}_P - \mathcal{A}_N\mathcal{T} + \mathcal{V} \\ 0 & \mathcal{A}_N \end{pmatrix}, \quad \text{where } \mathbf{T} = \begin{pmatrix} \mathbf{I} & \mathcal{T} \\ 0 & \mathbf{I} \end{pmatrix}$$

$\mathcal{T}$  is chosen as the solution to the sylvester equation

$$\mathcal{T}\mathcal{A}_P - \mathcal{A}_N\mathcal{T} + \mathcal{V} = 0$$

The original triple  $(\mathbf{A}, \mathbf{b}, \mathbf{c})$  is now given by  $(\mathbf{TQAQ}^{-1}\mathbf{T}^{-1}, \mathbf{TQb}, \mathbf{cQ}^{-1}\mathbf{T}^{-1})$

## Composition of Rational Functions

Working in the broader class of 2-EPT densities where the pointmass at zero is included we consider the composition of two rational functions,  $G$ , with minimal realization  $(\mathbf{A}_1, \mathbf{b}_1, \mathbf{c}_1, d_1)$  and,  $\phi$ , with minimal realization  $(\mathbf{A}_2, \mathbf{b}_2, \mathbf{c}_2, d_2)$ .

$$G(\phi(u)) = d_1 + \mathbf{c}_1 \left( \frac{1}{\phi(u)} - \mathbf{A}_1 \right)^{-1} \mathbf{b}_1$$

where

$$\begin{aligned} \phi(u) &= \mathbf{c}_2(u\mathbf{I} - \mathbf{A}_2)^{-1} \mathbf{b}_2 + d_2 \\ (\phi(u))^{-1} &= d_2^{-1} - d_2^{-1} \mathbf{c}_2 (u\mathbf{I} - \mathbf{A}_2 + \mathbf{b}_2 d_2^{-1} \mathbf{c}_2)^{-1} \mathbf{b}_2 d_2^{-1} \end{aligned}$$

Implementing Proposition (3.2) from Hanzon and Scherrer (2008) we have that  $G(\phi(u))$  is a rational function with realization

$$\begin{aligned} \mathbf{A} &= \mathbf{I}_{n_1} \otimes \tilde{\mathbf{A}}_2 + (\mathbf{A}_1 - \tilde{d}_2 \mathbf{I}_{n_1})^{-1} \otimes \tilde{\mathbf{b}}_2 \tilde{\mathbf{c}}_2 \\ \mathbf{b} &= -(\mathbf{A}_1 - \tilde{d}_2 \mathbf{I}_{n_1})^{-1} \mathbf{b}_1 \otimes \tilde{\mathbf{b}}_2 \\ \mathbf{c} &= \mathbf{c}_1 (\mathbf{A}_1 - \tilde{d}_2 \mathbf{I}_{n_1})^{-1} \otimes \tilde{\mathbf{c}}_2 \\ d &= d_1 - \mathbf{c}_1 (\mathbf{A}_1 - \tilde{d}_2 \mathbf{I}_{n_1})^{-1} \mathbf{b}_1 \end{aligned}$$

where  $\tilde{d}_2$  is not an eigenvalue of  $\mathbf{A}_1$

## A 2-EPT Approach to Financial Modelling with Variance Gamma

# Variance Gamma Density

$$f_{VG}(x; C, G, M) = \frac{(GM)^C}{\sqrt{\pi} \Gamma(C)} \exp\left(\frac{(G - M)x}{2}\right) \left(\frac{|x|}{G + M}\right)^{C-1/2} K_{C-1/2}\left(\frac{(G + M)|x|}{2}\right)$$

Variance Gamma (VG) Characteristic Function

$$\phi(ui; C, G, M) = \left( \frac{GM}{GM + (M - G)iu + u^2} \right)^C$$

For integer "C" the characteristic function is rational and the density simplifies to

$$f_{VG}(x; C, G, M) = \exp\left(\frac{(G - M)x - (G + M)|x|}{2}\right) \underbrace{\frac{(MG)^C}{(C - 1)!} \sum_{k=0}^{C-1} \frac{(C - 1 + k)!(G + M)^{-C-k} |x|^{C-1-k}}{(C - 1 - k)! k!}}_{p(x)}$$

# Variance Gamma Density

Splitting the density around the origin

$$f_{VG}(x; C, G, M) = \begin{cases} \exp(Gx) \frac{(MG)^C}{(C-1)!} \sum_{s=0}^{C-1} \frac{(2(C-1)-s)!(G+M)^{-2C+1+s}|x|^s}{s!(C-1-s)!} & \text{if } x \leq 0 \\ \exp(-Mx) \frac{(MG)^C}{(C-1)!} \sum_{s=0}^{C-1} \frac{(2(C-1)-s)!(G+M)^{-2C+1+s}|x|^s}{s!(C-1-s)!} & \text{if } x > 0 \end{cases}$$

Polynomial part equal for all  $x$  implying  $\mathbf{c}_P = \mathbf{c}_N = \mathbf{c}$  where entries are

$$c_s = \frac{(MG)^C}{(C-1)!} \frac{(2(C-1)-s)!(G+M)^{-2C+1+s}}{(C-1-s)!} \quad \dots \quad s \in \{0, 1, \dots, C-1\}$$

Similarly  $\mathbf{b}_P = \mathbf{b}_N = \mathbf{b} = (1, 0, \dots, 0)^T$  and let  $\mathbf{a}$  be a  $C \times C$  nilpotent matrix

$$f_{VG}(x; C, G, M) = \begin{cases} \mathbf{c} e^{Gx} e^{-\mathbf{a}x} \mathbf{b} = \mathbf{c} e^{(G\mathbf{I}-\mathbf{a})x} \mathbf{b} = \mathbf{c} e^{\mathbf{A}_N x} \mathbf{b} & \text{if } x \leq 0 \\ \mathbf{c} e^{-Mx} e^{\mathbf{a}x} \mathbf{b} = \mathbf{c} e^{(-M\mathbf{I}+\mathbf{a})x} \mathbf{b} = \mathbf{c} e^{\mathbf{A}_P x} \mathbf{b} & \text{if } x > 0 \end{cases}$$

# Variance Gamma Price Process

The Variance Gamma Process as described in Carr, Ghang, Madan (1998) under the risk neutral measure  $\mathbb{Q}$

$$S_T = S_t e^{(\omega+r)(T-t) + x_{T-t}}$$

To ensure  $\mathbb{E}_{\mathbb{Q}}[S_T] = S_t e^{r\tau}$ , where  $\tau = T - t$

$$\omega = -C \ln\left(\frac{MG}{MG + (M - G) - 1}\right)$$

Hence the shifted log return is given by

$$\ln\left(\frac{S_T}{S_t}\right) - (\omega + r)\tau = X_\tau \sim VG(C\tau, G, M)$$

# Vanilla Option Pricing and Greeks

A closed form expression can be obtained for European Call and Put Options

$$e^{-r\tau} \mathbb{E}_{\mathbb{Q}}[S(T) - K | S(T) > K] = -S(t)e^{\omega\tau} \mathbf{c}(\mathbf{A}_P + \mathbf{I})^{-1} e^{(\mathbf{A}_P + \mathbf{I}_C)d} \mathbf{b} - K e^{-r\tau} (1 - \Phi(d))$$

where

$$(1 - \Phi(d)) = \int_d^{\infty} \mathbf{c} e^{\mathbf{A}_P x} \mathbf{b} dx \quad , \quad d = \ln(K/S(t)) - (r + \omega)(\tau) > 0$$

Analytic solutions also exist for the Greeks and for  $d > 0$  we have

$$\frac{\partial C}{\partial S} = -e^{\omega\tau} \mathbf{c}(\mathbf{A}_P + \mathbf{I}_C)^{-1} e^{(\mathbf{A}_P + \mathbf{I}_C)d} \mathbf{b} + e^{\omega\tau} \mathbf{c} e^{(\mathbf{A}_P + \mathbf{I}_C)d} d - \frac{K}{S} e^{-r\tau} \mathbf{c} e^{\mathbf{A}_P d} \mathbf{b}$$

Value-at-Risk calculates a loss  $L$  for which a probability  $p$  exists of such a loss occurring

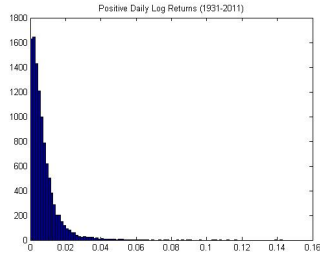
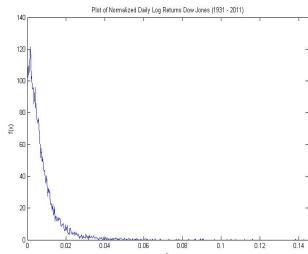
$$\begin{aligned} \mathbb{P}[S(T) < L] &= \mathbb{P}[S(0)e^{((r+\omega)T+X_T)} < L] &= p \\ &\mathbf{c}(\mathbf{A}_N)^{-1} e^{\mathbf{A}_N d_L} \mathbf{b} &= p \end{aligned}$$

where we solve for  $d_L = \ln\left(\frac{L}{S(0)}\right) - (r + \omega)T$



## Approximating Empirical Data with EPT Densities using RARL2

# Empirical Asset Returns



Problem: Derive the triple  $(\tilde{\mathbf{A}}_P, \tilde{\mathbf{b}}_P, \tilde{\mathbf{c}}_P)$  such that

$$\left( \log\left(\frac{S_{t+1}}{S_t}\right) \mid \frac{S_{t+1}}{S_t} > 1 \right) \sim EPT(\tilde{\mathbf{A}}_P, \tilde{\mathbf{b}}_P, \tilde{\mathbf{c}}_P)$$

where

$$\min_{\{\tilde{\mathbf{A}}, \tilde{\mathbf{b}}, \tilde{\mathbf{c}}\}} \|\tilde{\mathbf{c}} e^{\tilde{\mathbf{A}}x} \tilde{\mathbf{b}} - \tilde{f}(x)\|_2^2$$

# Sampled Data

Transform in continuous time given by rational function

$$\begin{aligned}\tilde{F}(s) &= \int_0^{\infty} \tilde{\mathbf{c}}_P e^{\tilde{\mathbf{A}}_P x} \tilde{\mathbf{b}}_P e^{-sx} dx = \tilde{\mathbf{c}}_P (s\mathbf{I} - \tilde{\mathbf{A}}_P)^{-1} \tilde{\mathbf{b}}_P \\ &= \frac{1}{N} \sum_{k=0}^N e^{-sx_k}\end{aligned}$$

However RARL2 requires samples to be inputted in discrete time  $F(z)$ ,  $F(e^{i\theta})$ . Isometry from Hanzon (1988) used to transform from continuous to discrete time

$$\tilde{F}(s) \mapsto F(z) = \frac{\sqrt{2}}{z-1} \tilde{F}\left(\frac{z+1}{z-1}\right) = \mathbf{c}_P (s\mathbf{I} - \mathbf{A}_P)^{-1} \mathbf{b}_P$$

where the triple  $(\mathbf{A}_P, \mathbf{b}_P, \mathbf{c}_P)$  is the discrete time realization.

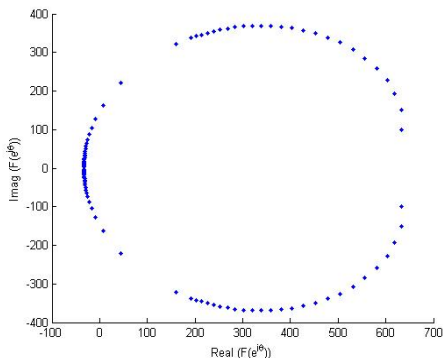
The Mobius transformation was employed to achieve this

$$z \mapsto s = \frac{z+1}{z-1}, \quad i\omega_j = \frac{e^{i\theta_j} + 1}{e^{i\theta_j} - 1}$$

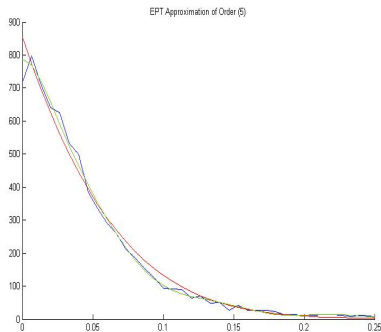
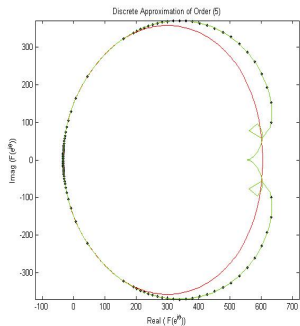
# Sampled Data

Hence we let  $\theta \in [-\pi, \pi]$ , solve for  $i\omega$  and find

$$F(e^{i\theta_j}) = \frac{\sqrt{2}}{e^{i\theta_j} - 1} \left( \frac{1}{N} \sum_{k=0}^N e^{-i\omega_j x_k} \right)$$



# Approximation with Sampled Data of Order 5



# Non-Negative EPT Densities

Suppose  $h(t)$  is the true function which we are approximating with the EPT density  $f(t) = \mathbf{c}e^{\mathbf{A}t}\mathbf{b}$  the minimization problem is

$$\min_{\{\mathbf{A}, \mathbf{b}, \mathbf{c} | f(t) \geq 0\}} \|h(t) - f(t)\|_2^2$$

Current RARL2 Algorithm: Minimize Eq. (5) w.r.t  $\mathbf{A}, \mathbf{c}$  and then choose optimal  $\mathbf{b}$

$$\min_{\{\mathbf{A}, \mathbf{c}\}} \|h(t) - f(t)\|_2^2 \quad (5)$$

Perron-Frobenius type result: *suppose  $f(t)$  is non-negative  $\forall t \geq 0$  then  $\lambda_M = \max_{\lambda \in \sigma(A)} \Re(\lambda)$  is an element of the spectrum of  $A$*

Hence a dominant real pole must be present in  $\mathbf{A}$ .

# Non-Negative EPT Densities

Possible Solution: Try to identify the dominant pole and its coefficient,  $\hat{\mu}e^{\hat{\lambda}_M t}$  from data or higher order system and exclude from approximation

$$\hat{h}(t) = h(t) - \hat{\mu}e^{\hat{\lambda}_M t}$$

$$\hat{f}(t) = f(t) - \hat{\mu}e^{\hat{\lambda}_M t} = \hat{\mathbf{c}}e^{\hat{\mathbf{A}}t}\hat{\mathbf{b}}$$

$$\min_{\{\hat{\mathbf{A}}, \hat{\mathbf{c}}\}} \|\hat{h}(t) - \hat{f}(t)\|_2^2 = \min_{\{\hat{\mathbf{A}}, \hat{\mathbf{c}}\}} \|\hat{h}(t) - \hat{\mathbf{c}}e^{\hat{\mathbf{A}}t}\hat{\mathbf{b}}\|_2^2$$

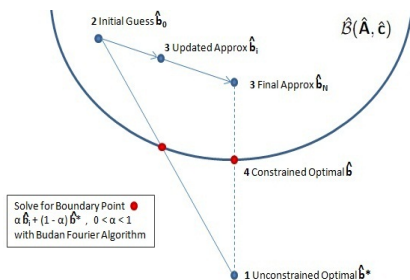
subject to  $\mathbb{R}(\sigma(\hat{\mathbf{A}})) < \hat{\lambda}_M$

$\hat{\mathcal{B}}(\hat{\mathbf{A}}, \hat{\mathbf{c}})$  is a convex set for a given  $(\hat{\mathbf{A}}, \hat{\mathbf{c}})$

Then there exists  $\hat{\mathbf{b}} \in \hat{\mathcal{B}}(\hat{\mathbf{A}}, \hat{\mathbf{c}})$  such that

$$\hat{f}(t) = \hat{\mathbf{c}}e^{\hat{\mathbf{A}}t}\hat{\mathbf{b}} \geq -\hat{\mu}e^{\hat{\lambda}_M t} \quad \forall t \geq 0$$

# Non-Negative EPT Densities



- 1 Find unconstrained optimal  $\hat{\mathbf{b}}^*$
- 2 If  $\hat{\mathbf{b}}^* \notin \hat{\mathcal{B}}(\hat{\mathbf{A}}, \hat{\mathbf{c}})$  then find initial estimate  $\hat{\mathbf{b}}_0 \in \hat{\mathcal{B}}(\hat{\mathbf{A}}, \hat{\mathbf{c}})$
- 3 Move from  $\hat{\mathbf{b}}_i$  to  $\hat{\mathbf{b}}_{i+1}$  based on  $L^2$  distance of boundary point (in red) to  $\hat{\mathbf{b}}^*$
- 4 Find constrained optimal  $\hat{\mathbf{b}} \in \hat{\mathcal{B}}(\hat{\mathbf{A}}, \hat{\mathbf{c}})$  s.t.  $\min_{\hat{\mathbf{b}}} \|\hat{\mathbf{b}} - \hat{\mathbf{b}}^*\|_2^2$

Using Budan Fourier Algorithm Hanzon (2010), we can locate boundary point for  $\hat{\mathbf{b}}_i$

It is a convex problem so we should be able to find constrained optimal



## Conclusions and Further Research

# Conclusions and Further Research

## Results

- Define 2-EPT Densities on whole real line
- Demonstrate various calculations with 2-EPT Densities
- Illustrate benefits of using 2-EPT Approach
- Approximating EPT Densities with RARL2

## Areas of Further Research with RARL2

- Approximate 2-EPT densities using RARL2 subject to the continuity condition  $\mathbf{c}_N \mathbf{b}_N = \mathbf{c}_P \mathbf{b}_P$
- Apply RARL2 to forward rate curves including Svensson and Nelson-Siegel curves e.g.  $y(t) = z_0 + z_1 e^{-\lambda_1 t} + z_2 t e^{-\lambda_2 t}$
- Constrain the approximating EPT functions to be non-negative

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# Thank You

Questions